

B8.3 Mathematical Models for Financial Derivatives

Hilary Term 2019

Solution Sheet Two

In the following $(W_t)_{t \geq 0}$ denotes a standard Brownian motion and $t \geq 0$ denotes time. A partition π of the interval $[0, t]$ is a sequence of points $0 = t_0 < t_1 < t_2 < \dots < t_n = t$ and $|\pi| = \max_k (t_{k+1} - t_k)$. On a given partition $W_k \equiv W_{t_k}$, $\delta W_k \equiv W_{k+1} - W_k$, $\delta t_k \equiv t_{k+1} - t_k$ and if f is a function on $[0, t]$, $f_k \equiv f(t_k)$ and $\delta f_k \equiv f_{k+1} - f_k$.

1. Show that if $t, s \geq 0$ then $\mathbb{E}[W_s W_t] = \min(s, t)$.

If $s = t \geq 0$ then we have

$$\mathbb{E}[W_s W_t] = \mathbb{E}[W_t^2] = t.$$

If not, assume $0 \leq s < t$ and write $W_t = W_s + W_t - W_s$. Then

$$\begin{aligned} \mathbb{E}[W_s W_t] &= \mathbb{E}[W_s W_s + W_s (W_t - W_s)] \\ &= \mathbb{E}[W_s W_s] + \mathbb{E}[W_s (W_t - W_s)] \\ &= \mathbb{E}[W_s^2] = s = \min(s, t). \end{aligned}$$

2. Suppose we define the following two stochastic integrals, the ‘anti-Itô’ integral

$$\int_0^t f(W_s, s) \bullet dW_s = \lim_{|\pi| \rightarrow 0} \sum_{k=0}^{n-1} f(W_{k+1}, t_{k+1}) (W_{k+1} - W_k),$$

and the Stratonovich integral

$$\begin{aligned} \int_0^t f(W_s, s) \circ dW_s \\ = \lim_{|\pi| \rightarrow 0} \sum_{k=0}^{n-1} \frac{1}{2} (f(W_{k+1}, t_{k+1}) + f(W_k, t_k)) (W_{k+1} - W_k). \end{aligned}$$

Show that

$$\begin{aligned} 2 \int_0^t W_s \bullet dW_s &= W_t^2 + t, & \mathbb{E} \left[2 \int_0^t W_s \bullet dW_s \right] &= 2t, \\ 2 \int_0^t W_s \circ dW_s &= W_t^2, & \mathbb{E} \left[2 \int_0^t W_s \circ dW_s \right] &= t \end{aligned}$$

$$\begin{aligned}
\int_0^t 2W_s \bullet dW_s &= \lim_{|\pi| \rightarrow 0} \sum_{k=0}^{n-1} 2W_{k+1} (W_{k+1} - W_k) \\
&= \lim_{|\pi| \rightarrow 0} \sum_{k=0}^{n-1} (W_{k+1}^2 - W_k^2) + \lim_{|\pi| \rightarrow 0} \sum_{k=0}^{n-1} (W_{k+1} - W_k)^2 \\
&= W_n^2 - W_0^2 + \lim_{|\pi| \rightarrow 0} \sum_{k=0}^{n-1} (W_{k+1} - W_k)^2 \\
&= W_t^2 + t.
\end{aligned}$$

Since we know $\mathbb{E}[W_t^2] = t$, we see that

$$\mathbb{E}\left[\int_0^t 2W_s \bullet dW_s\right] = \mathbb{E}[W_t^2] + t = 2t.$$

Similarly we have

$$\begin{aligned}
\int_0^t 2W_s \circ dW_s &= \lim_{|\pi| \rightarrow 0} \sum_{k=0}^{n-1} (W_{k+1} + W_k)(W_{k+1} - W_k) \\
&= \lim_{|\pi| \rightarrow 0} \sum_{k=0}^{n-1} (W_{k+1}^2 - W_k^2) \\
&= W_n^2 - W_0^2 \\
&= W_t^2,
\end{aligned}$$

and so

$$\mathbb{E}\left[\int_0^t 2W_s \circ dW_s\right] = \mathbb{E}[W_t^2] = t.$$

3. Assuming that both the integral and its variance exist, show that

$$\text{var}\left[\int_0^t f(W_s, s) dW_s\right] = \int_0^t \mathbb{E}[f(W_s, s)^2] ds.$$

[Note: if the integral and its variance exist then it is legitimate to interchange the order of expectation and integration.]

If the integral exists then its expected value is zero so

$$\text{var}\left[\int_0^t f(W_s, s) dW_s\right] = \mathbb{E}\left[\left(\int_0^t f(W_s, s) dW_s\right)^2\right].$$

Set $f_k = f(W_k, t_k)$ and consider

$$\begin{aligned}
\left(\sum_{k=0}^{n-1} f_k \delta W_k\right)^2 &= \sum_{k=0}^{n-1} \sum_{j=0}^{n-1} f_j f_k \delta W_j \delta W_k \\
&= \sum_{k=0}^{n-1} f_k^2 (\delta W_k)^2 + 2 \sum_{k=1}^{n-1} \sum_{j < k} f_j f_k \delta W_j \delta W_k.
\end{aligned}$$

We can eliminate the double sum using the tower law,

$$\begin{aligned}\mathbb{E}\left[f_j f_k \delta W_j \delta W_k\right] &= \mathbb{E}\left[\mathbb{E}_{t_k}\left[f_j f_k \delta W_j \delta W_k\right]\right] \\ &= \mathbb{E}\left[f_j f_k \delta W_j \mathbb{E}_{t_k}\left[\delta W_k\right]\right]\end{aligned}$$

since $t_j < t_k$, which means f_j and δW_j are known at time t_k , as is $f_k = f(W_k, t_k)$. As $\mathbb{E}_{t_k}[\delta W_k] = 0$, this term vanishes, leaving

$$\begin{aligned}\mathbb{E}\left[\left(\sum_{k=0}^{n-1} f_k \delta W_k\right)^2\right] &= \mathbb{E}\left[\sum_{k=0}^{n-1} f_k^2 (\delta W_k)^2\right] \\ &= \mathbb{E}\left[\sum_{k=0}^{n-1} \mathbb{E}_{t_k}\left[f_k^2 (\delta W_k)^2\right]\right] \\ &= \mathbb{E}\left[\sum_{k=0}^{n-1} f_k^2 \mathbb{E}_{t_k}\left[(\delta W_k)^2\right]\right] \\ &= \mathbb{E}\left[\sum_{k=0}^{n-1} f_k^2 \delta t_k\right] = \sum_{k=0}^{n-1} \mathbb{E}\left[f_k^2\right] \delta t_k\end{aligned}$$

and as we refine the partition, $|\pi| \rightarrow 0$, we see

$$\lim_{|\pi| \rightarrow 0} \sum_{k=0}^{n-1} \mathbb{E}\left[f_k^2\right] \delta t_k \rightarrow \int_0^t \mathbb{E}\left[f(W_s, s)^2\right] ds.$$

4. Use the differential version of Itô's lemma to show that

$$(a) \int_0^t W_s ds = t W_t - \int_0^t s dW_s = \int_0^t (t - s) dW_s,$$

We have

$$d(t W_t) = W_t dt + t dW_t$$

which integrates to show

$$t W_t = \int_0^t W_s ds + \int_0^t s dW_s.$$

Rearranging gives

$$\int_0^t W_s ds = t W_t - \int_0^t s dW_s = \int_0^t (t - s) dW_s.$$

$$(b) \int_0^t W_s^2 dW_s = \frac{1}{3} W_t^3 - \int_0^t W_s ds,$$

This time

$$d(W_t^3) = 3 W_t^2 dW_t + 3 W_t dt$$

which integrates to give

$$W_t^3 = 3 \int_0^t W_s^2 dW_s + 3 \int_0^t W_s ds.$$

Dividing by 3 and rearranging gives

$$\int_0^t W_s^2 dW_s = \frac{1}{3} W_t^3 - \int_0^t W_s ds.$$

$$(c) \mathbb{E}[e^{aW_t - a^2 t/2}] = 1.$$

Set $f(W, t) = \exp(aW - \frac{1}{2}a^2 t)$ and $f_t = f(W_t, t)$. Itô's lemma gives

$$\begin{aligned} df_t &= -\frac{1}{2} a^2 f_t dt + a f_t dW_t + \frac{1}{2} a^2 f_t dt \\ &= a f_t dW_t. \end{aligned}$$

This integrates to give $f_t - f_0 = a \int_0^t f_s dW_s$ or, in full,

$$\exp(aW_t - \frac{1}{2}a^2 t) = 1 + a \int_0^t \exp(aW_s - \frac{1}{2}a^2 s) dW_s$$

Taking expectations gives

$$\begin{aligned} \mathbb{E}\left[\exp(aW_t - \frac{1}{2}a^2 t)\right] &= 1 + a \mathbb{E}\left[\int_0^t \exp(aW_s - \frac{1}{2}a^2 s) dW_s\right] \\ &= 1. \end{aligned}$$

5. Define X_t to be the ‘area under a Brownian motion’, $X_0 = 0$ and $X_t = \int_0^t W_u du$ for $t > 0$. Show that X_t is normally distributed with

$$\mathbb{E}[X_t] = 0, \quad \mathbb{E}[X_t^2] = \frac{1}{3} t^3.$$

From Question 4(a) we have

$$\int_0^t W_s ds = \int_0^t (t-s) dW_s.$$

As $t - s$ does not depend on W_s , the integral is normally distributed with

$$\mathbb{E}\left[\int_0^t W_s ds\right] = \mathbb{E}\left[\int_0^t (t-s) dW_s\right] = 0$$

and

$$\begin{aligned}\text{var}\left[\int_0^t W_s ds\right] &= \text{var}\left[\int_0^t (t-s) dW_s\right] \\ &= \int_0^t (t-s)^2 ds \\ &= \frac{1}{3}t^3.\end{aligned}$$

Note that X_t is continuously differentiable for $t > 0$, with $\dot{X}_t = W_t$ (recall W_t is continuous in t).

Now define Y_t as the ‘average area under a Brownian motion’,

$$Y_t = \begin{cases} 0 & \text{if } t = 0, \\ X_t/t & \text{if } t > 0. \end{cases}$$

Show that Y_t has $\mathbb{E}[Y_t] = 0$, $\mathbb{E}[Y_t^2] = t/3$ and that Y_t is continuous for all $t \geq 0$.

Is $\sqrt{3}Y_t$ a Brownian motion? Give reasons for your answer.

For any $t > 0$, Y_t is normal because X_t is, with

$$\mathbb{E}[Y_t] = \mathbb{E}[X_t]/t = 0$$

and

$$\text{var}[Y_t] = \text{var}[X_t/t] = \text{var}[X_t]/t^2 = \frac{1}{3}t.$$

For $t > 0$ we have $Y_t = X_t/t$ which is the ratio of two differentiable functions and as the denominator is never zero it follows that Y_t is differentiable for $t > 0$, which implies continuous. Moreover, it means we can use l’Hopital’s rule to show

$$\lim_{t \rightarrow 0^+} Y_t = \lim_{t \rightarrow 0^+} \frac{W_t}{1} = 0,$$

which shows that Y_t is continuous at $t = 0$.

The function $\sqrt{3}Y_t$ is not only continuous but *differentiable* for $t > 0$ and therefore it can’t be a Brownian motion. (The hard way to do this part of the question is to show that the increments over disjoint intervals are not independent.)

6. Show that if

$$\frac{dS_t}{S_t} = \mu dt + \sigma dW_t,$$

and $f_t = f(S_t) = S_t^3$, then

$$\frac{df_t}{f_t} = 3(\mu + \sigma^2) dt + 3\sigma dW_t.$$

Write the SDE for S_t as

$$dS_t = \mu S_t dt + \sigma S_t dW_t$$

so Itô's lemma for a general $F_t = F(S_t, t)$ is

$$dF_t = \left(\frac{\partial F}{\partial t}(S_t, t) + \frac{1}{2} \sigma^2 S_t^2 \frac{\partial^2 F}{\partial S^2}(S_t, t) \right) dt + \frac{\partial F}{\partial S}(S_t, t) dS_t.$$

With $f(S) = S^3$, so $f'(S) = 3S^2$ and $f''(S) = 6S$, this becomes

$$\begin{aligned} df_t &= 3\sigma^2 S_t^3 dt + 3S_t^2 dS_t \\ &= 3 \left(\sigma^2 f_t dt + S_t^2 (\mu S_t dt + \sigma S_t dW_t) \right) \\ &= 3 f_t ((\mu + \sigma^2) dt + \sigma dW_t) \end{aligned}$$

which we could write as

$$\frac{df_t}{f_t} = 3(\mu + \sigma^2) dt + 3\sigma dW_t.$$

7. Find solutions of the Black-Scholes terminal value problem

$$\begin{aligned} \frac{\partial V}{\partial t} + \frac{1}{2} \sigma^2 S^2 \frac{\partial^2 V}{\partial S^2} + (r - y) S \frac{\partial V}{\partial S} - r V &= 0, \quad S > 0, \quad t < T, \\ V(S, T) &= f(S), \quad S > 0, \end{aligned}$$

when

(a) $f(S) = C$, where C is a constant;

In this case it makes sense to look for solutions which don't depend on S , so $\partial V / \partial S = 0$ and $\partial^2 V / \partial S^2 = 0$ identically. Then we have $V = V(t)$ and the partial differential equation reduces to the ODE

$$\frac{dV}{dt} - rV = 0, \quad V(T) = C$$

which has the solution $V = C e^{-r(T-t)}$.

(b) $f(S) = S^\alpha$, where α is a constant.

[Hint: you don't need the Feynman-Kac formula to do either of these. Look for simple functional forms of the solution.]

In this case, let's look for separable solutions

$$V(S, t) = S^\alpha f(t), \quad f(T) = 1.$$

We find that

$$\frac{\partial V}{\partial t} = S^\alpha \dot{f}(t), \quad S \frac{\partial V}{\partial S} = \alpha S^\alpha f(t), \quad S^2 \frac{\partial^2 V}{\partial S^2} = \alpha(\alpha - 1) S^\alpha f(t)$$

and the Black-Scholes equation becomes

$$S^\alpha \left(\dot{f}(t) + \frac{1}{2} \sigma^2 \alpha(\alpha - 1) f(t) + (r - y) \alpha f(t) - r f(t) \right) = 0.$$

As this has to hold for all $S > 0$ we have an ordinary differential equation for $f(t)$ which we can write as

$$\dot{f}(t) = \lambda f(t), \quad f(T) = 1,$$

where

$$\lambda = r - (r - y) \alpha - \frac{1}{2} \sigma^2 \alpha(\alpha - 1).$$

The solution is

$$f(t) = e^{-\lambda(T-t)}$$

and

$$V(S, t) = e^{-\lambda(T-t)} S^\alpha.$$

Optional questions

8. The Black-Scholes equation from a binomial method.

One step of the Cox, Ross & Rubinstein binomial method can be written as

$$V(S, t) = e^{-r\delta t} \left(q V^u + (1 - q) V^d \right)$$

where

$$\begin{aligned} V^u &= V(S^u, t + \delta t), \quad V^d = V(S^d, t + \delta t), \\ S^u &= S e^{\sigma\sqrt{\delta t}}, \quad S^d = S e^{-\sigma\sqrt{\delta t}}, \quad q = \frac{e^{r\delta t} - e^{-\sigma\sqrt{\delta t}}}{e^{\sigma\sqrt{\delta t}} - e^{-\sigma\sqrt{\delta t}}}, \end{aligned}$$

$\sigma > 0$ is the volatility, r is the risk-free rate and $\delta t > 0$ is the length of the time-step. Supposing this is true for all $S > 0$ and that $V(S, t)$ may be expanded in a Taylor series in both S and t , show that as $\delta t \rightarrow 0$

$$\begin{aligned} q &= \frac{1}{2} + \frac{r - \frac{1}{2}\sigma^2}{2\sigma} \sqrt{\delta t} + \mathcal{O}(\delta t), \\ V^u &= V + \sqrt{\delta t} \sigma S \frac{\partial V}{\partial S} + \delta t \left(\frac{\partial V}{\partial t} + \frac{1}{2}\sigma^2 \left(S \frac{\partial V}{\partial S} + S^2 \frac{\partial^2 V}{\partial S^2} \right) \right) + \mathcal{O}(\delta t^{3/2}), \\ V^d &= V - \sqrt{\delta t} \sigma S \frac{\partial V}{\partial S} + \delta t \left(\frac{\partial V}{\partial t} + \frac{1}{2}\sigma^2 \left(S \frac{\partial V}{\partial S} + S^2 \frac{\partial^2 V}{\partial S^2} \right) \right) + \mathcal{O}(\delta t^{3/2}), \end{aligned}$$

where V and all its partial derivatives are evaluated at (S, t) . Hence show that in the limit $\delta t \rightarrow 0$ the option price satisfies the (zero dividend-yield) Black-Scholes equation

$$\frac{\partial V}{\partial t} + \frac{1}{2}\sigma^2 S^2 \frac{\partial^2 V}{\partial S^2} + r S \frac{\partial V}{\partial S} - r V = 0.$$

As $\delta t \rightarrow 0$ we can expand

$$\begin{aligned} e^{\sigma\sqrt{\delta t}} &= 1 + \sigma\sqrt{\delta t} + \frac{1}{2}\sigma^2\delta t + \mathcal{O}(\delta t^{3/2}), \\ e^{-\sigma\sqrt{\delta t}} &= 1 - \sigma\sqrt{\delta t} + \frac{1}{2}\sigma^2\delta t + \mathcal{O}(\delta t^{3/2}), \\ e^{r\delta t} &= 1 + r\delta t + \mathcal{O}(\delta t^2), \end{aligned}$$

from which it follows that

$$\begin{aligned} q &= \frac{\sigma\sqrt{\delta t} + (r - \frac{1}{2}\sigma^2)\delta t + \mathcal{O}(\delta t^{3/2})}{2\sigma\sqrt{\delta t} + \mathcal{O}(\delta t^{3/2})} \\ &= \frac{1}{2} + \frac{r - \frac{1}{2}\sigma^2}{2\sigma} \sqrt{\delta t} + \mathcal{O}(\delta t), \end{aligned}$$

and that

$$\begin{aligned}
& V(S e^{\sigma\sqrt{\delta t}}, t + \delta t) \\
&= V(S, t) + \delta t \frac{\partial V}{\partial t}(S, t) + (\sigma\sqrt{\delta t} + \tfrac{1}{2}\sigma^2\delta t) S \frac{\partial V}{\partial S}(S, t) \\
&\quad + \tfrac{1}{2}\sigma^2\delta t S^2 \frac{\partial^2 V}{\partial S^2}(S, t) + \mathcal{O}(\delta t^{3/2}) \\
&= V + \sqrt{\delta t} \sigma S \frac{\partial V}{\partial S} + \delta t \left(\frac{\partial V}{\partial t} + \tfrac{1}{2}\sigma^2 \left(S \frac{\partial V}{\partial S} + S^2 \frac{\partial^2 V}{\partial S^2} \right) \right) + \mathcal{O}(\delta t^{3/2}) \\
& V(S e^{-\sigma\sqrt{\delta t}}, t + \delta t) \\
&= V - \sqrt{\delta t} \sigma S \frac{\partial V}{\partial S} + \delta t \left(\frac{\partial V}{\partial t} + \tfrac{1}{2}\sigma^2 \left(S \frac{\partial V}{\partial S} + S^2 \frac{\partial^2 V}{\partial S^2} \right) \right) + \mathcal{O}(\delta t^{3/2}).
\end{aligned}$$

Therefore

$$\begin{aligned}
& q V(S e^{\sigma\sqrt{\delta t}}, t + \delta t) + (1 - q) V(S e^{-\sigma\sqrt{\delta t}}, t + \delta t) \\
&= V + \delta t \left(\frac{\partial V}{\partial t} + \tfrac{1}{2}\sigma^2 S^2 \frac{\partial^2 V}{\partial S^2} + \tfrac{1}{2}\sigma^2 S \frac{\partial V}{\partial S} \right) + (r - \tfrac{1}{2}\sigma^2) \delta t S \frac{\partial V}{\partial S} + \mathcal{O}(\delta t^{3/2}) \\
&= V + \delta t \left(\frac{\partial V}{\partial t} + \tfrac{1}{2}\sigma^2 S^2 \frac{\partial^2 V}{\partial S^2} + r S \frac{\partial V}{\partial S} \right) + \mathcal{O}(\delta t^{3/2})
\end{aligned}$$

and hence

$$\begin{aligned}
& e^{-r\delta t} \left(q V(S e^{\sigma\sqrt{\delta t}}, t + \delta t) + (1 - q) V(S e^{-\sigma\sqrt{\delta t}}, t + \delta t) \right) \\
&= (1 - r\delta t) \left(V + \delta t \left(\frac{\partial V}{\partial t} + \tfrac{1}{2}\sigma^2 S^2 \frac{\partial^2 V}{\partial S^2} + r S \frac{\partial V}{\partial S} \right) \right) + \mathcal{O}(\delta t^{3/2}) \\
&= V + \delta t \left(\frac{\partial V}{\partial t} + \tfrac{1}{2}\sigma^2 S^2 \frac{\partial^2 V}{\partial S^2} + r S \frac{\partial V}{\partial S} - r V \right) + \mathcal{O}(\delta t^{3/2}).
\end{aligned}$$

Substituting this back into the binomial equation gives

$$V = V + \delta t \left(\frac{\partial V}{\partial t} + \tfrac{1}{2}\sigma^2 S^2 \frac{\partial^2 V}{\partial S^2} + r S \frac{\partial V}{\partial S} - r V \right) + \mathcal{O}(\delta t^{3/2})$$

and cancelling V and dividing by δt then gives

$$\left(\frac{\partial V}{\partial t} + \tfrac{1}{2}\sigma^2 S^2 \frac{\partial^2 V}{\partial S^2} + r S \frac{\partial V}{\partial S} - r V \right) + \mathcal{O}(\delta t^{1/2}) = 0.$$

Taking the limit $\delta t \rightarrow 0$ gives the Black-Scholes equation,

$$\frac{\partial V}{\partial t} + \tfrac{1}{2}\sigma^2 S^2 \frac{\partial^2 V}{\partial S^2} + r S \frac{\partial V}{\partial S} - r V = 0.$$

9. The variation of a function, or stochastic process, over $[0, t]$, is

$$\langle f \rangle_t = \lim_{|\pi| \rightarrow 0} \sum_{k=0}^{n-1} |f_{k+1} - f_k|.$$

If $\langle f \rangle_t$ is finite on $[0, t]$ we say f has bounded variation on $[0, t]$. Show that:

(a) if f is $C^1[0, t]$ then $\langle f \rangle_t = \int_0^t |f'(s)| ds < \infty$;

If f is C^1 on $[0, t]$ then the intermediate value theorem asserts that $f_{k+1} - f_k = f'(\xi_k)(t_{k+1} - t_k)$ for some $\xi_k \in [t_k, t_{k+1}]$. Since $\delta t_k = (t_{k+1} - t_k) > 0$,

$$\langle f \rangle_t = \lim_{|\pi| \rightarrow 0} \sum_{k=0}^{n-1} |f'(\xi_k)| \delta t_k = \int_0^t |f'(s)| ds.$$

If $f'(s)$ is continuous on $[0, t]$ then so too is $|f'(s)|$ and as $[0, t]$ is a closed and bounded set, $|f'(s)|$ must be bounded on $[0, t]$. Therefore the integral must be finite and hence f has bounded variation on $[0, t]$.

(b) if f is a continuous function with $\langle f \rangle_t < \infty$ then its quadratic variation is zero, $[f]_t = 0$;

We have

$$[f]_t = \lim_{|\pi| \rightarrow 0} \sum_{k=0}^{n-1} (f_{k+1} - f_k)^2 \geq 0$$

and we also have

$$\sum_{k=0}^{n-1} (f_{k+1} - f_k)^2 \leq \left(\max_j |f_{j+1} - f_j| \right) \left(\sum_{k=0}^{n-1} |f_{k+1} - f_k| \right).$$

If f has bounded variation then the sum on the right-hand side is finite as $|\pi| \rightarrow 0$. If f is continuous then the maximum value of $|f_{j+1} - f_j| \rightarrow 0$ as $|\pi| \rightarrow 0$. Thus the product of the two vanishes as $|\pi| \rightarrow 0$.

(c) Brownian motion does not have bounded variation;

Brownian motion is continuous in t and has non-zero quadratic variation. Therefore it must have unbounded variation by the previous part of the question.

(d) the arc length of the graph of a Brownian motion is infinite for any $t > 0$.

[Hint: if $y = X_t$ has an arc length s then $ds = \sqrt{dy^2 + dt^2} \geq \sqrt{dy^2} = |dy|$.]

Assume that we can define arc length for a Brownian motion. Then it is given by

$$s = \int_0^t ds$$

where if $y = X_t$ then ds is given by

$$ds^2 = dt^2 + dy^2, \quad ds = \sqrt{dt^2 + dy^2}.$$

Now observe that

$$ds = \sqrt{dt^2 + dy^2} \geq \sqrt{dy^2} = |dy|.$$

If the integral of $|dy|$ existed it would equal the variation of W_t , but we know that W_t doesn't have bounded variation. This implies that the length of the graph of a Brownian motion is infinite (for any $t > 0$).