

B8.3 Week 3 Summary 2019

Brownian motion

A stochastic process is a sequence of random variables indexed by a parameter, for example, $(W_t)_{t \geq 0}$. For each fixed $t \geq 0$, W_t is a random variable.

A process $(W_t)_{t \geq 0}$ is a Brownian motion if (and only if)

1. $\forall s \geq 0, t \geq 0$, $(W_{t+s} - W_t)$ is normally distributed with zero mean and variance s ,

$$\mathbb{E}[W_{t+s} - W_t] = 0, \quad \mathbb{E}[(W_{t+s} - W_t)^2] = s,$$

2. if $0 \leq p \leq q \leq s \leq t$ then $(W_q - W_p)$ and $(W_t - W_s)$ are independent,
3. the map $t \mapsto W_t$ is continuous, and
4. $W_0 = 0$ (this is really a convention, it saves some writing).

It is not obvious that such a thing exists, but there are a number of ways of constructing it (see Etheridge §3.1 and §3.2, for example).

Note that if $(W_t)_{t \geq 0}$ is a Brownian motion then so too are:

1. $\hat{W}_t = W_{(t+t_0)} - W_{t_0}$ for any constant $t_0 \geq 0$;
2. $\tilde{W}_t = c W_{(t/c^2)}$ for any constant $c > 0$.

Brownian motion is almost surely not differentiable

We show that with probability one a Brownian motion is not *differentiable*. If Brownian motion were differentiable at the point $t_0 \geq 0$ then the limit

$$\lim_{t \rightarrow 0} \frac{W_{(t+t_0)} - W_{t_0}}{t} = \lim_{t \rightarrow 0} \frac{\hat{W}_t}{t}$$

would exist, so it is enough to show that with probability one the second limit does exist. Let A_n and B_n be defined by

$$A_n = \left\{ \frac{|\hat{W}_t|}{t} > n : \text{for some } t \in \left(0, \frac{1}{n^4}\right] \right\}, \quad B_n = \left\{ \frac{|\hat{W}_t|}{t} > n : \text{at } t = \frac{1}{n^4} \right\}.$$

Clearly we have $B_n \subseteq A_n$ and so

$$\begin{aligned} \text{prob}(A_n) &\geq \text{prob}(B_n) = \text{prob}\left(\frac{|\hat{W}_{1/n^4}|}{1/n^4} > n\right) \\ &= \text{prob}\left(|n^2 \hat{W}_{1/n^4}| > \frac{1}{n}\right) \\ &= \text{prob}\left(|\tilde{W}_1| > \frac{1}{n}\right). \end{aligned}$$

As $n \rightarrow \infty$ we have $\text{prob}(|\tilde{W}_1| > 1/n) \rightarrow 1$. Therefore $\lim_{n \rightarrow \infty} \text{prob}(A_n) = 1$ which means that in this limit there is (with probability one) always some $0 < t \leq 1/n^4$ with $|\hat{W}_t|/t > n$. This shows that (with probability one) the limit which defines the derivative of a Brownian motion can not exist.

Quadratic variation

Let π be a partition of $[0, t]$,

$$t_0 = 0 < t_1 < t_2 < \cdots < t_n = t,$$

and let

$$|\pi| = \max_{0 \leq k < n} (t_{k+1} - t_k).$$

The quadratic variation of a function f on $[0, t] \rightarrow \mathbb{R}$ is defined to be¹

$$[f]_t = \lim_{|\pi| \rightarrow 0} \sum_{k=0}^{n-1} (f_{k+1} - f_k)^2$$

where $f_k = f(t_k)$. It may or may not exist, depending on f .

1. If f is continuously differentiable on $[0, t]$ then $[f]_t = 0$.

As $f_{k+1} - f_k = f'(\xi_k)(t_{k+1} - t_k)$ for some $\xi_k \in [t_k, t_{k+1}]$ we have

$$\begin{aligned} \sum_{k=0}^{n-1} (f_{k+1} - f_k)^2 &= \sum_{k=0}^{n-1} f'(\xi_k)^2 (t_{k+1} - t_k)^2 \\ &\leq |\pi| \sum_{k=0}^{n-1} f'(\xi_k)^2 (t_{k+1} - t_k) \end{aligned}$$

and as $|\pi| \rightarrow 0$

$$\sum_{k=0}^{n-1} f'(\xi_k)^2 (t_{k+1} - t_k) \rightarrow \int_0^t f'(u)^2 du < \infty,$$

using Riemann's definition of an integral (which is equivalent to Lebesgue's definition if the function is continuous, as it is in this case).

2. The quadratic variation of a Brownian motion is defined as

$$[W]_t = \lim_{|\pi| \rightarrow 0} \sum_{k=0}^{n-1} (W_{k+1} - W_k)^2,$$

¹Other common notation for quadratic variation include $[f]_t^2$ and $[f, f]_t$.

where $W_k = W_{t_k}$. We find that $[W]_t = t$, in the sense that

$$\mathbb{E}[[W]_t - t] = 0, \quad \mathbb{E}[([W]_t - t)^2] = 0.$$

It follows that Brownian motion is almost surely *not* continuously differentiable in t .

To see this, let $\delta W_k = W_{k+1} - W_k$ and $\delta t_k = t_{k+1} - t_k$

$$\mathbb{E}\left[\sum_{k=0}^{n-1} ((\delta W_k)^2 - \delta t_k)\right] = \sum_{k=0}^{n-1} (\mathbb{E}[(\delta W_k)^2] - \delta t_k),$$

which vanishes for any finite $n > 0$ since $\mathbb{E}[(\delta W_k)^2] = \delta t_k$. It therefore also vanishes in the limit $n \rightarrow \infty$.

Next, consider

$$\begin{aligned} & \mathbb{E}\left[\left(\sum_{k=0}^{n-1} ((\delta W_k)^2 - \delta t_k)\right)^2\right] \\ &= \sum_{j=0}^{n-1} \sum_{k=0}^{n-1} \mathbb{E}[(\delta W_j)^2 - \delta t_j][(\delta W_k)^2 - \delta t_k] \\ &= \sum_{k=0}^{n-1} \mathbb{E}[(\delta W_k)^2 - \delta t_k]^2 + \sum_{j \neq k}^{n-1} \mathbb{E}[(\delta W_j)^2 - \delta t_j] \mathbb{E}[(\delta W_k)^2 - \delta t_k] \\ &= \sum_{k=0}^{n-1} \mathbb{E}[(\delta W_k)^2 - \delta t_k]^2 \\ &= \sum_{k=0}^{n-1} (\mathbb{E}[(\delta W_k)^4] - 2\delta t_k \mathbb{E}[(\delta W_k)^2] + (\delta t_k)^2) \\ &= 2 \sum_{k=0}^{n-1} (\delta t_k)^2 \\ &\leq 2|\pi| \sum_{k=0}^{n-1} \delta t_k = 2|\pi|t, \end{aligned}$$

where we use the independence of δW_j and δW_k if $j \neq k$ to get from the second to the third line.

The Itô integral

The definition of the Itô integral of a function against a Brownian motion is

$$\int_0^t f(W_u, u) dW_u = \lim_{|\pi| \rightarrow 0} \sum_{k=0}^{n-1} f(W_k, t_k)(W_{k+1} - W_k).$$

For fixed t this integral is a random variable and as t varies it is a stochastic process. The sum converges to the integral in an L^2 sense (see, e.g., Etheridge pp 78–85).

Using the tower law, and writing $\delta W_k = W_{k+1} - W_k$, we find that

$$\begin{aligned}\mathbb{E}\left[\sum_{k=0}^{n-1} f(W_k, t_k) \delta W_k\right] &= \mathbb{E}\left[\sum_{k=0}^{n-1} \mathbb{E}_{t_k}[f(W_k, t_k) \delta W_k]\right] \\ &= \mathbb{E}\left[\sum_{k=0}^{n-1} f(W_k, t_k) \mathbb{E}_{t_k}[\delta W_k]\right] = 0.\end{aligned}$$

This establishes that if the Itô integral exists then

$$\mathbb{E}\left[\int_0^t f(W_u, u) dW_u\right] = 0.$$

The same sort of argument shows that for $0 \leq s < t$

$$\mathbb{E}_s\left[\int_0^t f(W_u, u) dW_u\right] = \int_0^s f(W_u, u) dW_u.$$

If f is a reasonable function of t alone then

$$\begin{aligned}\mathbb{E}\left[\left(\sum_{k=0}^{n-1} f(t_k) \delta W_k\right)^2\right] &= \mathbb{E}\left[\sum_{j,k=0}^{n-1} f(t_j) f(t_k) \delta W_k \delta W_j\right] \\ &= \mathbb{E}\left[\sum_{k=0}^{n-1} \mathbb{E}_k[f(t_k)^2 (\delta W_k)^2]\right] + 2\mathbb{E}\left[\sum_{\substack{j < k \\ j,k=0 \\ j,k=0}}^{n-1} \mathbb{E}_k[f(t_j) f(t_k) \delta W_j \delta W_k]\right] \\ &= \mathbb{E}\left[\sum_{k=0}^{n-1} f(t_k)^2 \mathbb{E}_k[(\delta W_k)^2]\right] + 2\mathbb{E}\left[\sum_{\substack{j < k \\ j,k=0 \\ j,k=0}}^{n-1} f(t_j) f(t_k) \delta W_j \mathbb{E}_k[\delta W_k]\right] \\ &= \mathbb{E}\left[\sum_{k=0}^{n-1} f(t_k)^2 \mathbb{E}_{t_k}[(\delta W_k)^2]\right] \\ &= \sum_{k=0}^{n-1} f(t_k)^2 \delta t_k\end{aligned}$$

and in the limit $|\pi| \rightarrow 0$ we obtain *Itô's isometry*,

$$\text{var}\left[\int_0^t f(u) dW_u\right] = \int_0^t f(u)^2 du.$$

As the integral is simply the limit of a sum of normally distributed random variables, it is itself normally distributed (proof omitted).

If f depends on W_t the same sort of argument shows that

$$\text{var} \left[\int_0^t f(W_u, u) dW_u \right] = \int_0^t \mathbb{E}[f(W_u, u)^2] du,$$

provided the right-hand side exists. In general, however, the integral itself is not normally distributed. For example, $2 \int_0^t W_u dW_u = W_t^2 - t$, which has a χ^2 distribution.

Itô's lemma

If $f(x, \tau)$ is $C^{2,1}$ then

$$f(W_t, t) - f(0, 0) =$$

$$\int_0^t \frac{\partial f}{\partial \tau}(W_u, u) du + \int_0^t \frac{\partial f}{\partial x}(W_u, u) dW_u + \frac{1}{2} \int_0^t \frac{\partial^2 f}{\partial x^2}(W_u, u) d[W]_u. \quad (1)$$

Since $[W]_u = u$ we can replace $d[W]_u$ by du , and in practice we always do.

Consider the simpler case where f is independent of τ and write

$$f(W_t) - f(0) = \sum_{k=0}^{n-1} (f(W_{k+1}) - f(W_k))$$

over some partition, π , of $[0, t]$. Taylor's theorem (with remainders) shows that for each k

$$f(W_{k+1}) - f(W_k) = f'(W_k) \delta W_k + \frac{1}{2} f''(V_k) (\delta W_k)^2$$

for some V_k between W_k and W_{k+1} , where $\delta W_k = W_{k+1} - W_k$. Thus

$$f(W_t) - f(0) = \sum_{k=0}^{n-1} f'(W_k) \delta W_k + \frac{1}{2} \sum_{k=0}^{n-1} f''(V_k) (\delta W_k)^2.$$

As we refine the partition

$$\lim_{|\pi| \rightarrow 0} \sum_{k=0}^{n-1} f'(W_k) \delta W_k \rightarrow \int_0^t f'(W_u) dW_u.$$

For the second sum, it can be shown that

$$\lim_{|\pi| \rightarrow 0} \sum_{k=0}^{n-1} f''(V_k) (\delta W_k)^2 \rightarrow \int_0^t f''(W_u) d[W]_u,$$

establishing that

$$f(W_t) - f(0) = \int_0^t f'(W_u) dW_u + \frac{1}{2} \int_0^t f''(W_u) d[W]_u.$$

Itô's lemma in practice

In practice, we usually write (1) in differential form rather than an integral form. If $f(W, t)$ is $C^{2,1}$ and we define $f_t = f(W_t, t)$ the differential form of Itô's lemma is

$$df_t = \left(\frac{\partial f}{\partial t}(W_t, t) + \frac{1}{2} \frac{\partial^2 f}{\partial W^2}(W_t, t) \right) dt + \frac{\partial f}{\partial W}(W_t, t) dW_t.$$

This amounts to doing a regular Taylor series expansion of $f(W, t)$ then *pretending* that $dW_t^2 = dt$ (and ignoring terms of higher order than dt).

To solve the stochastic differential equation

$$\frac{dS_t}{S_t} = \mu dt + \sigma dW_t \quad (2)$$

we can proceed as follows. If $f(W, t) = e^{aW+bt}$ then all its partial derivatives are multiples of the function, so it makes sense to try

$$S_t = S_0 e^{aW_t+bt}.$$

This gives

$$dS_t = (b S_t + \frac{1}{2} a^2 S_t) dt + a S_t dW_t$$

or

$$\frac{dS_t}{S_t} = (b + \frac{1}{2} a^2) dt + a dW_t.$$

If we set $a = \sigma$ and $b = \mu - \frac{1}{2} \sigma^2$ we recover (2), i.e., the solution of (2) is

$$S_t = S_0 \exp\left(\left(\mu - \frac{1}{2} \sigma^2\right) t + \sigma W_t\right).$$

The process S_t is often called *geometric* Brownian motion. Note that the sign of S_t is determined by the sign of S_0 .

Itô's lemma for solutions of SDEs

Suppose that X_t is a solution of

$$X_t - X_0 = \int_0^t \mu(X_u, u) du + \int_0^t \sigma(X_u, u) dW_u,$$

$f(x, t)$ is a $C^{2,1}$ function and we set $f_t = f(X_t, t)$. Then

$$f_t - f_0 = \int_0^t \left(\frac{\partial f}{\partial t}(X_u, u) + \frac{1}{2} \sigma(X_u, u)^2 \frac{\partial^2 f}{\partial x^2}(X_u, u) \right) du + \int_0^t \frac{\partial f}{\partial x}(X_u, u) dX_u$$

The proof amounts to showing the quadratic variation $[X]_t$ is given by

$$[X]_t = \int_0^t \sigma(X_u, u)^2 du.$$

In differential notation, which is how this result is normally used, if

$$dX_t = \mu(X_t, t) dt + \sigma(X_t, t) dW_t \quad (3)$$

and $f_t = f(X_t, t)$ then

$$df_t = \left(\frac{\partial f}{\partial t}(X_t, t) + \frac{1}{2} \sigma(X_t, t)^2 \frac{\partial^2 f}{\partial x^2}(X_t, t) \right) dt + \frac{\partial f}{\partial x}(X_t, t) dX_t. \quad (4)$$

This can be obtained from a Taylor series expansion of $f(x, t)$ and *pretending* that $dX_t^2 = \sigma(X_t, t)^2 dt$.

The Feynman K ac formula

Suppose that $f(x, t)$ satisfies the terminal value problem

$$\begin{aligned} \frac{\partial f}{\partial t} + \frac{1}{2} \sigma(x, t)^2 \frac{\partial^2 f}{\partial x^2} + \mu(x, t) \frac{\partial f}{\partial x} &= 0, \quad t < T, \quad x \in \mathbb{R}, \\ f(x, T) &= F(x), \quad x \in \mathbb{R}. \end{aligned} \quad (5)$$

Let X_t satisfy the stochastic differential equation

$$dX_t = \mu(X_t, t) dt + \sigma(X_t, t) dW_t$$

Then

$$f(x, t) = \mathbb{E}_t[F(X_T) \mid X_t = x] \quad (6)$$

To see this, note that It 's lemma implies that

$$\begin{aligned} f(X_T, T) &= f(X_t, t) + \int_t^T \sigma(X_s, s) \frac{\partial f}{\partial x}(X_s, s) dW_s \\ &\quad + \int_t^T \left(\frac{\partial f}{\partial t}(X_s, s) + \mu(X_s, s) \frac{\partial f}{\partial x}(X_s, s) + \frac{1}{2} \sigma(X_s, s)^2 \frac{\partial^2 f}{\partial x^2}(X_s, s) \right) ds \end{aligned}$$

By assumption, the integral on the second line vanishes and when we take expectations the integral on the first line also vanishes. Thus

$$f(X_t, t) = \mathbb{E}_t[f(X_T, T)]$$

and conditioning on $X_t = x$ gives

$$f(x, t) = \mathbb{E}_t[f(X_T, T) \mid X_t = x]$$